



Variance Estimation Using Linear Combination of Skewness and Quartiles

M. A. Bhat^{1*}, S. Maqbool², S. A. Saraf², Ab. Rouf² and S. H. Malik²

¹Division of Agricultural Economics and Statistics, SKUAST-Kashmir, India.

²Division of Economics and Statistics, SKUAST-Kashmir, India.

Authors' contributions

This work was carried out in collaboration between all authors. Author MAB designed the study, performed the statistical analysis, wrote the protocol and wrote the first draft of the manuscript. Author SM managed the analysis of the study. Authors SAS, AR and SHM managed the literature searches. All authors read and approved the final manuscript.

Article Information

DOI: 10.9734/AIR/2018/37321

Editor(s):

(1) Dawit Getnet Ayele, Department of Epidemiology, The Johns Hopkins University, Bloomberg School of Public Health, USA

Reviewers:

(1) Davidson Odafe Akpootu, Usmanu Danfodiyo University Sokoto, Nigeria.

(2) Kunio Takezawa, Institute for Agro-Environmental Sciences, Japan.

Complete Peer review History: <http://www.sciencedomain.org/review-history/22761>

Original Research Article

Received 10th October 2017
Accepted 14th November 2017
Published 19th January 2018

ABSTRACT

In this paper we have suggested Modified ratio type variance estimators where in our aim is to estimate the population variance in the presence of outliers, when there is strong correlation between auxiliary variable and study variable by using, the linear combination of skewness and quartiles as auxiliary information. To judge the efficiency of suggested estimators over existing estimators practically, we have carried out the Bias and Mean square error of proposed and existing estimators and suggested estimators have proven better performance than the existing estimators.

Keywords: Simple random sampling; bias; mean square error; skewness and quartiles efficiency.

1. INTRODUCTION

Here we consider a finite population $U = \{U_1, U_2, \dots, U_N\}$ of N distinct and

identifiable units. Let Y be a real variable with value Y_i measured on $U_i, i = 1, 2, 3, \dots, N$ given a vector $[Y_1, Y_2, Y_3, \dots, Y_N]$. Sometimes in sample

*Corresponding author: E-mail: mabhat.1500@gmail.com;

surveys information on auxiliary variable X correlated with study variable Y, is available can be utilized to obtain the efficient estimator for the estimation of Population variance. To estimate the population variance, various efficient estimators have been widely discussed by the authors such as Isaki [1] who proposed ratio and regression estimators. Latter various authors such as Kadilar & Cingi, H [2], have also proposed the ratio estimators to improve the efficiency of modified estimators over existing estimators. Subramani, J. and Kumarapandiyam, G. [3], Kadilar and Cingi, H. [4], have also contributed a lot to the theory of ratio type variance estimation. Similarly, authors such as

Arcos, A., M. Rueda, M. D. Martinez, S. Gonzalez and Y. Roman. [5], who incorporated the auxiliary information available in variance estimation.

Jeelani, Iqbal and Maqbool, S. [6], proposed the modified ratio type estimators for population mean by using linear combination of coefficient of skewness and quartile deviation as auxiliary variables to enhance the efficiency of proposed estimators.

Kadilar, C., and Cingi, H, 2004 [7], suggested new ratio type estimators using correlation coefficient as auxiliary variable. Accordingly, various authors such as, Upadhyaya and Singh [8], Murthy M. N [9] and Singh D, and Chaudhary, F. S. [10] has also modified various estimators to improve the efficiency of existing estimators. "In this paper our aim is to estimate the population variance

random sample selected from population U" in the presence of outliers.

2. MATERIALS AND METHODS

2.1 Notations

N = Population size. n = Sample size. $\gamma = \frac{1}{n}$
 Y = study variable. X = Auxiliary variable. \bar{X}, \bar{Y} = Population means. \bar{x}, \bar{y} = Sample means.
 S_y^2, S_x^2 = population variances. s_y^2, s_x^2 = sample variances. C_x, C_y = Coefficient of variation

ρ = coefficient $\beta_{1(x)}$ = Skewness of auxiliary variable $\beta_{2(x)}$ = Kurtosis of the auxiliary variable.

$\beta_{2(y)}$ = Kurtosis of the study variable. M_d = Median of the auxiliary variable. $B(.)$ = Bias of the estimator. $MSE(.)$ = Mean square error. \hat{S}_R^2 = Ratio type variance estimator.

\hat{S}_{US1}^2 = Existing estimator proposed by Upadhyaya & Singh

\hat{S}_{Kc1}^2 = Existing modified ratio estimator (proposed by Kadilar & Cingi), \hat{S}_{JG}^2 = Existing Modified ratio estimator (proposed by J.Subramani & G. Kumarapandian), Q_d = Quartile deviation, Q_a = Quartile average, Q_1 = first quartile, Q_2 = second quartile Q_3 = third quartile $Q_r = Q_3 - Q_1$ quartile range. $\lambda_{rs} = \frac{\mu_{rs}}{\mu_{20}^{r/2} \mu_{02}^{s/2}}$

Where $\mu_{rs} = \frac{1}{N} \sum (Y_i - \bar{Y})^r (X_i - \bar{X})^s$

In this paper, we discuss the already existing estimators in the literature and then proposed modified estimators where we have used the linear combination of skewness and quartiles and accordingly we will compare the results of suggested estimators with the existing estimators.

2.2 Existing Estimators

2.2.1 Ratio type Variance estimator proposed by Upadhyaya and Singh [8]

Isaki suggested a ratio type variance estimator for the population variance S_y^2 when the population variance S_x^2 of the auxiliary variable X is known. Bias and mean square error is given by

$$\hat{S}_{US1}^2 = s_y^2 \left[\frac{S_x^2 + \beta_{2x}}{s_x^2 + \beta_{2x}} \right] \quad (1)$$

$$\text{Bias}((\hat{S}_R^2) = \gamma S_y^2 \left[(\beta_{2(x)} - 1) - (\lambda_{22} - 1) \right] \quad (2)$$

$$\text{MSE}((\hat{S}_R^2) = \gamma S_y^4 \left[(\beta_{2(y)} - 1) + (\beta_{2(x)} - 1) - 2(\lambda_{22} - 1) \right] \quad (3)$$

2.2.2 Ratio type variance estimator proposed by Kadilar and Cingi [2]

The authors have suggested ratio type variance estimators where they have used known values of Coefficient of variance and coefficient of kurtosis as an auxiliary variable X, whose Bias, mean square and percent relative efficiency has been shown in Table-1, Table-2, Table-3, Table-4 and Table-5

$$\hat{S}_{kc1}^2 = S_y^2 \left[\frac{S_x^2 + C_x}{s_x^2 + C_x} \right] \tag{4}$$

$$\text{Bias} ((\hat{S}_{kc1}^2)) = \gamma S_y^2 A_1 [A_1(\beta_{2(x)} - 1) - (\lambda_{22} - 1)] \tag{5}$$

$$\text{MSE} ((\hat{S}_{kc1}^2)) = \gamma S_y^4 [(\beta_{2(y)} - 1) + A_1^2 (\beta_{2(x)} - 1) - 2A_1(\lambda_{22} - 1)] \tag{6}$$

2.2.3 Ratio type variance estimator proposed by J. Subramani & G. Kumara pandiyan [3]

The authors have suggested ratio type variance estimators and they have used median, quartiles and Deciles as an auxiliary variable X, whose bias, mean square error and percent relative efficiency is given in Table-1, Table-2, Table-3, Table-4 and Table-5 respectively

$$\hat{S}_{jG}^2 = S_y^2 \left[\frac{S_x^2 + \alpha w_i}{s_x^2 + \alpha w_i} \right] \tag{7}$$

$$\text{Bias} ((\hat{S}_{jG}^2)) = \gamma S_y^2 A_{jG} [A_{jG}(\beta_{2(x)} - 1) - (\lambda_{22} - 1)]$$

$$\text{Mse} ((\hat{S}_{jG}^2))$$

$$\gamma S_y^4 [(\beta_{2(y)} - 1) + A_{jG}^2 (\beta_{2(x)} - 1) - 2A_{jG}(\lambda_{22} - 1)] \tag{8}$$

2.3 Proposed Estimator

We have proposed a new modified ratio type variance estimator of the auxiliary variable by using linear combination of skewness and quartiles. Since quartiles are not sensitive to outliers, as they divide the series into different series and provide various location parameters and shape shifts, accounts the distributional properties and also estimates the covariate effects of average value

$$\begin{aligned} \hat{S}_{MS1}^2 &= S_y^2 \left[\frac{S_x^2 + (\beta_1 + Q_1)}{s_x^2 + (\beta_1 + Q_1)} \right] & \hat{S}_{MS2}^2 &= S_y^2 \left[\frac{S_x^2 + (\beta_1 + Q_2)}{s_x^2 + (\beta_1 + Q_2)} \right] \\ \hat{S}_{MS3}^2 &= S_y^2 \left[\frac{S_x^2 + (\beta_1 + Q_3)}{s_x^2 + (\beta_1 + Q_3)} \right] & \hat{S}_{MS4}^2 &= S_y^2 \left[\frac{S_x^2 + (\beta_1 + Q_d)}{s_x^2 + (\beta_1 + Q_d)} \right] \\ \hat{S}_{MS5}^2 &= S_y^2 \left[\frac{S_x^2 + (\beta_1 + Q_a)}{s_x^2 + (\beta_1 + Q_a)} \right] & \hat{S}_{MS6}^2 &= S_y^2 \left[\frac{S_x^2 + (\beta_1 + Q_r)}{s_x^2 + (\beta_1 + Q_r)} \right] \end{aligned}$$

Here we have derived the bias and mean square error of the proposed estimator \hat{S}_{MSi}^2 ; $i = 1, 2, \dots, 6$ up to the first order of approximation as given below:

$$\text{Let } e_0 = \frac{s_y^2 - S_y^2}{S_y^2} \text{ and } e_1 = \frac{s_x^2 - S_x^2}{S_x^2}. \text{ Further we}$$

can write $s_y^2 = S_y^2(1 + e_0)$ and $s_x^2 = S_x^2(1 + e_1)$ and from the definition of e_0 and e_1 we obtain:

$$\begin{aligned} E[e_0] &= E[e_1] = 0, & E[e_0^2] &= \frac{1-f}{n}(\beta_{2(y)} - 1), \\ E[e_1^2] &= \frac{1-f}{n}(\beta_{2(x)} - 1), & E[e_0 e_1] &= \frac{1-f}{n}(\lambda_{22} - 1) \end{aligned}$$

The proposed estimator \hat{S}_{MSi}^2 ; $i = 1, 2, \dots, 6$ is given below:

$$\hat{S}_{MSi}^2 = S_y^2 \left[\frac{S_x^2 + \alpha a_i}{s_x^2 + \alpha a_i} \right] \tag{1}$$

$$\Rightarrow \hat{S}_{MSi}^2 = S_y^2(1 + e_0) \left[\frac{S_x^2 + \alpha a_i}{s_x^2 + e_1 S_x^2 + \alpha a_i} \right]$$

$$\Rightarrow \hat{S}_{MSi}^2 = \frac{S_y^2(1 + e_0)}{(1 + A_{MSi} e_1)}$$

$$\text{Where } A_{MSi} = \frac{S_x^2}{S_x^2 + \alpha a_i}$$

is a constant and MS_i , $i = 1, 2, \dots, 6$ suggested estimators and $a_i = (\beta_1 + Q_i)$; $i = 1, 2, 3, d, a, r$

$$\Rightarrow \hat{S}_{MSi}^2 = S_y^2(1 + e_0)(1 + A_{MSi} e_1)^{-1} \tag{2}$$

$$\begin{aligned} \Rightarrow \hat{S}_{MSi}^2 &= S_y^2(1 + e_0)(1 - A_{MSi} e_1 + A_{MSi}^2 e_1^2 \\ &- A_{MSi}^3 e_1^3 + \dots) \end{aligned} \tag{3}$$

Table 1. Bias and MSE of existing estimators

Existing estimator	Population 1		Population 2		Population 3	
	Bias	Mean square error	Bias	Mean square error	Bias	Mean square error
Upadhyaya & Singh [8]	6.88	2678.64	189.61	8285277.52	267.48	8662043.54
Kadilar&Cingi [2]	7.51	2806.20	193.45	8305040.92	274.87	8700066.36
Subramani & Kumarapandiyam [3]	6.22	2551.09	194.29	8309384.71	276.27	8707285.44

Expanding and neglecting the terms more than 3rd order, we get

$$\hat{S}_{MSI}^2 = S_y^2 + S_y^2 e_0 - S_y^2 A_{MSI} e_1 - S_y^2 A_{MSI}^2 e_0 e_1 + S_y^2 A_{MSI}^2 e_1^2 \quad (4)$$

$$\Rightarrow \hat{S}_{MSI}^2 - S_y^2 = S_y^2 e_0 - S_y^2 A_{MSI} e_1 - S_y^2 A_{MSI}^2 e_0 e_1 + S_y^2 A_{MSI}^2 e_1^2 \quad (5)$$

By taking expectation on both sides of (5), we get

$$\begin{aligned} E(\hat{S}_{MSI}^2 - S_y^2) &= S_y^2 E(e_0) - S_y^2 A_{MSI} E(e_1) \\ &\quad - S_y^2 A_{MSI}^2 E(e_0 e_1) + S_y^2 A_{MSI}^2 E(e_1^2) \\ Bias(\hat{S}_{MSI}^2) &= S_y^2 A_{MSI}^2 E(e_1^2) - S_y^2 A_{MSI} E(e_0 e_1) \\ Bias(\hat{S}_{MSI}^2) &= \gamma S_y^2 A_{MSI} [A_{MSI} (\beta_{2(x)} - 1) - (\lambda_{22} - 1)] \quad (6) \end{aligned}$$

Squaring both sides of (5) and (6), neglecting the terms more than 2nd order and taking expectation, we get

$$\begin{aligned} E(\hat{S}_{MSI}^2 - S_y^2)^2 &= S_y^4 E(e_0^2) \\ &\quad + S_y^4 A_{MSI}^2 E(e_1^2) - 2S_y^4 A_{MSI} E(e_0 e_1) \\ MSE(\hat{S}_{MSI}^2) &= \gamma S_y^4 [(\beta_{2(y)} - 1) \\ &\quad + A_{MSI}^2 (\beta_{2(x)} - 1) - 2A_{MSI} (\lambda_{22} - 1)] \end{aligned}$$

3. RESULTS AND DATA ANALYSIS

3.1 Numerical Illustration

3.1.1 Population-1

We use the data of Murthy [9] page 228 in which fixed capital is denoted by X(auxiliary variable) and output of 80 factories are denoted by

Population-3 Singh & Chaudhary [10]

$$\begin{aligned} N = 34, n = 20, \bar{Y} = 85.64, \bar{X} = 19.94, S_y = 73.31, \lambda_{22} = 1.2244, S_x = 15.02, C_x = 0.7532, \beta_{2x} = 3.7257, \\ \beta_{2y} = 13.3666, \beta_{1x} = 1.2758, Q_1 = 9.925, Q_2 = 14.25, Q_3 = 27.8, Q_r = 17.87, Q_a = 18.86, Q_d = 8.9375 \end{aligned}$$

Y(study variable).we apply the proposed and existing estimators to this data set and the data statistics is given below:

$$\begin{aligned} N = 80, n = 20, S_x = 8.4563, C_x = 0.7507, \\ S_y = 18.3569, \bar{Y} = 51.8264, \beta_{1x} = 1.05 \\ \beta_{2x} = 2.8664, \beta_{2y} = 2.2667, \lambda_{22} = 2.2209, \\ Q_1 = 9.31, Q_2 = 7.5750, Q_3 = 16.9750, \\ Q_r = 11.82, Q_a = 11.0625, Q_d = 5.9125, \\ \bar{X} = 11.2624 \end{aligned}$$

3.1.2 Population-2

We have taken population-2 from Singh & Chaudhary (1986) given in page 177

Population-2: Singh and Chaudhary [10]

$$\begin{aligned} N = 34, n = 20, \bar{Y} = 85.64, \bar{X} = 20.88, \\ S_y = 73.31 \\ S_x = 15.05, \beta_{1x} = 0.8732, \beta_{2x} = 2.91, \beta_{2y} = 13.36, \\ \lambda_{22} = 1.1525, Q_1 = 9.42, Q_2 = 15.0, Q_3 = 25.47, \\ Q_r = 16.05, Q_a = 17.45, Q_d = 16.05, C_x = 0.7205 \end{aligned}$$

3.1.3 Population-3

We have taken population-3 again from Singh & Chaudhary (1986) given in page 177.

*Corresponding author: E-mail: mabhat.1500@gmail.com;

Table 2. Bias and MSE of proposed estimators

Existing estimator	Population-1		Population-2		Population-3	
	Bias	Mean square error	Bias	Mean square error	Bias	Mean square error
Upadhyaya & Singh [8]	6.88	2678.64	189.61	8285277.52	267.48	8662043.54
Kadilar & Cingi [2]	7.51	2806.20	193.45	8305040.92	274.87	8700066.36
Subramani & Kumarapandiyan [3]	6.22	2551.09	194.29	8309384.71	276.27	8707285.44
MS ₁	5.91	2508.57	176.24	817622.78	248.59	852774.58
MS ₂	5.32	2381.01	168.05	813275.44	238.83	847345.97
MS ₃	3.20	2125.91	153.51	805873.95	213.32	834887.79
MS ₄	4.31	2253.46	165.96	812979.38	231.92	844302.48
MS ₅	4.46	2254.46	164.49	811795.14	229.82	843064.95
MS ₆	5.75	2466.05	165.96	812929.38	250.23	853302.69

Table 3. Percent relative efficiency of proposed estimators with existing estimators for Population-1

Estimators	P1	P2	P3	P4	P5	P6
Upadhyaya & Singh [8]	106.77	112.50	125.99	118.86	118.86	108.62
Kadilar & Cingi [2]	111.86	117.85	131.99	124.52	124.52	113.79
Subramani & Kumarapandiyan [3]	101.69	107.14	119.99	113.20	113.20	103.44

Table 4. Percent relative efficiency of proposed estimators with existing estimators for Population-2

Estimators	P1	P2	P3	P4	P5	P6
Upadhyaya & Singh [8]	1013.33	1018.75	1028.11	1019.12	1020.61	1019.18
Kadilar & Cingi [2]	1015.75	1021.18	1030.56	1021.55	1023.04	1021.61
Subramani & Kumarapandiyan [3]	1016.28	1021.71	1031.10	1022.09	1023.58	1022.15

Table 5. Percent relative efficiency of proposed estimators with existing estimators for Population-3

Estimators	P1	P2	P3	P4	P5	P6
Upadhyaya & Singh [8]	1015.74	1022.25	1037.51	1025.94	1027.44	1015.69
Kadilar & Cingi [2]	1020.20	1026.74	1042.06	1030.44	1031.95	1019.57
Subramani & Kumarapandiyan [3]	1021.05	1027.59	1042.92	1031.29	1032.81	1020.42

4. CONCLUSION AND LIMITATION

In this paper, the above tables have clearly revealed that our proposed estimators are more efficient than the existing estimators when the comparison is made between the existing and proposed estimators together with their percent relative efficiency criteria. Hence the proposed estimator may be preferred over existing estimators for use in practical applications. Furthermore, the advantage of using these methods to estimate the population variance is that they are not sensitive to outliers, but with an disadvantage that they can be further modified to improve the efficiency of the variance estimators.

COMPETING INTERESTS

Authors have declared that no competing interests exist.

REFERENCES

1. Isaki CT. Variance estimation using auxiliary information. Journal of the American Statistical Association. 1983; 78:117-123.
2. Kadilar C, Cingi H. Improvement in variance estimation using auxiliary information. Hacettepe Journal of mathematics and Statistics. 2006; 35(1):117-115.

3. Subramani J, Kumarapandiyam G. Generalized modified ratio type estimator for estimation of population variance. Sri-Lankan Journal of Applied Statistics. 2015;16-1:69-90.
4. Kadilar, Cingi H. Ratio estimators for population variance in simple and stratified sampling. Applied Mathematics and computation. 2006b;173:1047-1058.
5. Arcos A, Rueda M, Martinez MD, Gonzalez S, Roman Y. Incorporating the auxiliary information available in variance estimation. Applied Mathematical and Computation. 2005;160:387-399.
6. Jeelani, Iqbal, Maqbool S. Modified ratio estimators of population mean using linear combination of coefficient of skewness and quartile deviation. The South Pacific Journal of Natural and Applied Sciences. 2013;31:39-44.
7. Kadilar C, Cingi H. New ratio estimators using correlation coefficient. Interstat. 2004;4:1-11.
8. Upadhyaya LN, Singh HP. Use of auxiliary variable in the estimation of population variance. Mathematical forum, 1999; 1936:4:33-36.
9. Murthy MN. Sampling theory. Theory and methods. Statistical Publishing Society, Calcutta; 1967.
10. Singh D, Chaudhary FS. Theory and analysis of sample survey designs. New Age Publishers; 1986.

© 2018 Bhat et al.; This is an Open Access article distributed under the terms of the Creative Commons Attribution License (<http://creativecommons.org/licenses/by/4.0>), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Peer-review history:

*The peer review history for this paper can be accessed here:
<http://www.sciencedomain.org/review-history/22761>*